

Branching type processes

and DTNs

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Dec 2009

Background on Branching

- 19th century: concern among Victorians about possible extinction of aristocratic surnames.
- Galton posed this question in the *Educational Times* of 1873. The Reverend Watson replied with a solution. Joint publication of the solution in 1874.
- The G-W process: $X_{n+1} = \sum_{i=1}^{X_n} \xi_n^{(i)}$.
- The G-W process with immigration: $X_{n+1} = \sum_{i=1}^{X_n} \xi_n^{(i)} + B_n$.
- Multitype: Y_n is a vector, $Y_{n+1} = A_n(Y_n) + B_n$, where

$$[A_n(Y_n)]_i = \sum_{j=1}^N \sum_{k=1}^{Y_n^j} \xi_{ji}^{(k)}(n)$$

Example 1: discrete branching with migration

Queue with Vacations, Gated Regime

- $M/G/1/\infty$ queue,
- Arrival rate λ , i.i.d. service times $\{D_n\}$ with first and second moments $d, d^{(2)}$.
- Sequence of vacations: V_n . Will be assumed stationary ergodic, with first and second moments $v, v^{(2)}$.
- Gated regime: at the n th end of vacation, a gate is closed (n th polling instant). Then the server goes on serving the customers present at the queue at that polling instant:
Then the server leaves on vacation.

• We denote:

- $B_n :=$ the number of arrivals during the n th vacation.
- $\xi_h^{(i)} :=$ the number of arrivals during the service time of a customer

• Then:

$$X_{n+1} = \sum_{i=1}^{X_n} \xi_n^{(i)} + B_n, \quad n \geq n_0.$$

• **Divisibility property:** Denote

$$A_n(x) = \sum_{i=1}^x \xi_n^{(i)}$$

Then A_n are nonnegative and divisible:

$$A_n(x + y) = A_n^{(1)}(x) + A_n^{(2)}(y)$$

where $A_n^{(i)}$ are i.i.d.

We use this property to define branching process. Applies for non-integers!

Example 2: continuous branching with migration

Queue with Vacations, Gated Regime

- Define the time to serve N customers as:

$$\tau(N) := \sum_{i=1}^N D_i$$

- Let $\mathcal{N}(T)$ denote the number of arrivals during a random duration T , where the arrival process is Poisson with rate λ , and is independent of T .
- Denote by $\hat{\mathcal{A}}_n(C_n) = \tau(\mathcal{N}(C_n))$, i.e. the sum of service times of all the arrivals during C_n .
- We obtain

$$C_{n+1} = \hat{\mathcal{A}}_n(C_n) + V_{n+1}. \quad (1)$$

Application: Message Ferry [Kavitha + EA]

- A bus has a given fixed trajectory on the plain.
- It starts at a BS (base station), collects packets. It has d stops. At each stop it transmits packets to the closest sensors.
- V_n is total walking time in cycle n . Previous equation holds for cycle time.
- Known as a globally gated polling system.

Queueing with gated vacations $d = 1$

Expected waiting time: (FIFO queue)

Consider an arbitrary customer. Upon arrival, it has to wait for

1. The residual cycle time C_{res} ,
2. The service time of all the customers that arrived during C_{past} which is the past cycle time: $d(\lambda E[C_{past}]) = \rho E[C_{past}]$

We have from [Baccelli & Brémaud, 1994]

$$E[C_{res}] = E[C_{past}] = \frac{E[C_0^2]}{2E[C_0]}.$$

Thus the expected waiting time of an arbitrary customer is given by

$$E[W_n] = (1 + \rho) \frac{E[C_0^2]}{2E[C_0]},$$

The expected number of customers in queue in stationary regime (not including service) is obtained using Little's Theorem: $\lambda E[W_n]$.

Conclusion: we need to compute $E[C_0]$ and $E[C_0^2]$!

1 Steady State Probabilities of CBP

Iterating $Y_{n+1} = A_n(Y_n) + B_n$, we obtain from A1:

$$\begin{aligned}
 Y_2 &= A_1(Y_1) + B_1 \\
 &= A_1(A_0(Y_0) + B_0) + B_1 \\
 &= A_1^{(0)}(A_0(Y_0)) + A_1^{(1)}(B_0) + B_1 \\
 &= A_1^{(0)} A_0^{(0)}(Y_0) + A_1^{(1)}(B_0) + B_1.
 \end{aligned}$$

$$\begin{aligned}
 Y_3 &= A_2(Y_2) + B_2 \\
 &= A_2(A_1(Y_1) + B_1) + B_2 \\
 &= A_2(A_1(A_0(Y_0) + B_0) + B_1) + B_2 \\
 &= A_2^{(0)} A_1^{(0)} A_0^{(0)}(Y_0) + A_2^{(1)} A_1^{(1)}(B_0) + A_2^{(2)}(B_1) + B_2
 \end{aligned}$$

$$Y_n = \sum_{j=0}^{n-1} \left(\prod_{i=n-j}^{n-1} A_i^{(n-j)} \right) (B_{n-j-1}) + \left(\prod_{i=0}^{n-1} A_i^{(0)} \right) (Y_0), \quad n > 0 \quad (2)$$

(we understand $\prod_{i=n}^k A_i(x) = x$ whenever $k < n$, and $\prod_{i=n}^k A_i(x) = A_k A_{k-1} \dots A_n$ whenever $k > n$).

- Under fairly general assumptions, $\lim_{n \rightarrow \infty} \left(\prod_{i=0}^{n-1} A_i^{(0)} \right) (y) = 0$, so Y_n has a limit as $n \rightarrow \infty$ distributed like

$$Y_n^* =_d \sum_{j=0}^{\infty} \left(\prod_{i=n-j}^{n-1} A_i^{(n-j)} \right) (B_{n-j-1}), \quad n \in \mathbb{Z}, \quad (3)$$

where for each integer i , $\{A_i^{(j)}(\cdot)\}_j$ have the same distribution as $A_i(\cdot)$.

- Sufficient condition: stationarity plus $\|\mathcal{A}\| < 1$.
- Branching processes: $\{A_i^{(j)}(\cdot)\}_j$ are i.i.d.
- Stochastic differential equations: they are equal.
- The representation holds for general dependence: **Semi linear processes**.

Moments:

- (i) The first moment of X_n^* is given by

$$E[X_0^*] = (I - \mathcal{A})^{-1}b. \quad (4)$$

- (ii) Assume that the first and second moments b_i and $b_i^{(2)}$'s are finite and that F satisfies

$$\lim_{n \rightarrow \infty} F^n = 0. \quad (5)$$

Define Q to be the matrix whose ij th entry is $Q_{ij} = \sum_{k=1}^d \bar{y}_k \Gamma^{(k)}$. Then the matrix $\text{cov}(X^*)$ is the unique solution of the set of linear equations:

$$\text{cov}(X) = \text{cov}(B) + \sum_{r=1}^{\infty} \left(\mathcal{A}^r \hat{\mathcal{B}}(r) + \left[\mathcal{A}^r \hat{\mathcal{B}}(r) \right]^T \right) + F(\text{cov}[X]) + Q. \quad (6)$$

The second moment matrix $E[XX^T]$ in steady state is the unique solution of the set of linear equations:

$$E[XX^T] = E[B_0 B_0^T] + \sum_{r=1}^{\infty} \left(\mathcal{A}^r \mathcal{B}(r) + \left[\mathcal{A}^r \mathcal{B}(r) \right]^T \right) + F(E[XX^T]) + Q. \quad (7)$$

2 Example: Delay Tolerant Ad-hoc Networks

- Delay tolerant Ad-hoc Networks make use of nodes' mobility to compensate for lack of instantaneous connectivity.
- Information sent by a source to a disconnected destination can be forwarded and relayed by other mobile nodes.
- Let X_n^+ be the number of nodes that have a copy of the packet at time n ,
- Let X_n^- be the number of nodes that do not have a copy of the packet at time n .
- Mobility: a mobile present at time n may leave and other may arrive. Let B_n be the number of new arrivals.

- Let $\rho_n^{(i)}$ and $\hat{\rho}_n^{(i)}$ be the indicator that node i remains in the system for the next slot. ρ is used for nodes that have the packet and $\hat{\rho}$ for the others.
- Let $\xi_n^{(i)}$ be the indicator that the source meets mobile i at time slot n . These are i.i.d. Then

$$X_{n+1}^+ = \sum_{i=1}^{X_n^+} \rho_n^{(i)} + \sum_{i=1}^{X_n^-} \hat{\rho}_n^{(i)} \xi_n^{(i)}$$

$$X_{n+1}^- = \sum_{i=1}^{X_n^-} \hat{\rho}_n^{(i)} (1 - \xi_n^{(i)}) + B_n$$

Controlling the Energy

- Assume that the source limits the transmissions in order to save energy
- Let ζ_n be the indicator that the source intends to transmit a packet at time n . Assume ζ_n are i.i.d.

$$X_{n+1}^+ = \sum_{i=1}^{X_n^+} \rho_n^{(i)} + \zeta_n \sum_{i=1}^{X_n^-} \hat{\rho}_n^{(i)} \xi_n^{(i)}$$

$$X_{n+1}^- = \sum_{i=1}^{X_n^-} \hat{\rho}_n^{(i)} (1 - \zeta_n \xi_n^{(i)}) + B_n$$

- This is a semi-linear process, not a branching process

Partial Information

- Observations: Assume that each node is "sampled" each time unit with some small probability.
- Let $\alpha_n^{(i)}$ be the indicator that node i is sampled at time n .

$$X_{n+1}^+ = \sum_{i=1}^{X_n^+} \rho_n^{(i)} + \sum_{i=1}^{X_n^-} \hat{\rho}_n^{(i)} \xi_n^{(i)}$$

$$X_{n+1}^- = \sum_{i=1}^{X_n^-} \hat{\rho}_n^{(i)} (1 - \xi_n^{(i)}) + B_n$$

$$Y_{n+1} = \sum_{i=1}^{X_n^+} \alpha_n^{(i)}$$

Filtering

- Objective: monitoring the number of packets.
- Example: first order linear filter
- Let \hat{X}_n^+ be the estimator of X_n^+ , The estimation error is $\epsilon_n = \hat{X}_n^+ - X_n^+$.

$$X_{n+1}^+ = \sum_{i=1}^{X_n^+} \rho_n^{(i)} + \sum_{i=1}^{X_n^-} \hat{\rho}_n^{(i)} \xi_n^{(i)}$$

$$X_{n+1}^- = \sum_{i=1}^{X_n^-} \hat{\rho}_n^{(i)} (1 - \xi_n^{(i)}) + B_n$$

$$Y_{n+1} = \sum_{i=1}^{X_n^+} \alpha_n^{(i)}$$

$$\hat{X}_{n+1} = K_1(Y_n - \hat{X}_n) + K_2 \hat{X}_n$$

$$\epsilon_n = \hat{X}_n - X_n$$

- Semi-linear process. We can compute $E[\epsilon_n^2]$ and compute K that minimizes it.